

# Recent and future developments in earthquake ground motion estimation

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## Abstract

Seismic hazard analyses (SHA) are routinely carried out around the world to understand the hazard, and consequently the risk, posed by earthquake activity. Whether single scenario, deterministic analyses, or state-of-the art probabilistic approaches, considering all possible events, a founding pillar of SHA is the estimation of the ground-shaking field from potential future earthquakes. Early models accounted for simple observations, such that ground shaking from larger earthquakes is stronger and that ground motion tends to attenuate rapidly away from the earthquake source. The first ground motion prediction equations (GMPEs) were, therefore, developed with as few as two principal predictor variables: magnitude and distance.

Despite the significant growth of computer power over the last few decades, and with it the possibility to compute kinematic or dynamic rupture models coupled with simulations of 3D wave propagation, the simple parametric GMPE has remained the tool of choice for hazard analysts. There are numerous reasons for this. First and foremost GMPEs are robust and reliable

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within the model space considered during their derivation, and many can be extrapolated to a degree beyond this space with some confidence. With ever expanding datasets and improved metadata the models are becoming more and more useful: a range of predictor variables are now used, describing the source, path and site effects in detail. GMPEs are also relatively easy to implement and computationally inexpensive. Despite this, probabilistic hazard calculations using GMPEs and accounting for uncertainties can still take several days to run. Full simulation-based approaches, therefore, clearly lie outside the computation budget afforded to most projects.

As well as the ever expanding list of predictor variables, other recent developments have also significantly improved the predictive power of GMPEs. This has allowed them to maintain their advantage over more ‘physical’ simulation techniques. Possibly the biggest aspect of this is not related to the median ground-shaking field, but rather its variability (and correlation in space and with oscillator period). This is a major advantage of empirical as opposed to simulation approaches, which typically struggle to replicate the covariance of input variables and, consequently, the variance of the ground motion. In this article we summarize some of the recent advances in ground motion prediction equations, including their application in SHA. We begin with a summary of the current state-of-the-art, then introduce the main additional predictor variables now used. Region- and event-type (tectonic or induced) specific predictions and adjustments are then discussed. Additional topics include advances in estimating ground-motion variability (epistemic and aleatory) and expanding GMPEs to predict other intensity measures or waveform features. The article concludes with a discussion on the path forward in earthquake ground motion prediction.

*Keywords:* seismology, earthquake engineering, earthquake, induced



## 1. Introduction

Seismic hazard assessment for a given site is founded on two pillars: firstly, a seismic-source model quantitatively describing all possible earthquakes in the vicinity (generally within about 300 km) and, secondly, a ground-motion model expressing the shaking that would happen at the site given the occurrence of each of these earthquakes. This article focuses on the second of these components; nevertheless, when considering ground-motion models it is vital to bear in mind the descriptions of earthquakes contained within the seismic-source model. These descriptions invariably consist of the earthquake's geographical location (and depth), its magnitude and, increasingly, its faulting mechanism and other characteristics (e.g. rupture geometry).

The results of seismic hazard assessments are vital inputs to earthquake engineering as they provide the motions that need to be resisted by structures and infrastructure constructed at the site. In the past most earthquake engineering analyses were based on the response spectral representation of shaking (e.g. Newmark and Hall, 1982; Chopra, 1995) or other pseudo-static methods. Consequently only estimates of scalar intensity measures (IMs), the principal ones being peak ground acceleration (PGA) and velocity (PGV) and elastic response spectral accelerations (SA) at various structural periods between 0 and commonly 2 s, were required for engineering analysis. In the past decade or so, Incremental Dynamic Analysis (Vamvatsikos and Cornell, 2002) and other time-history-based approaches have become increasingly used. There is a growing need, therefore, for seismic hazard

analysts to provide a time-history representation of earthquake shaking in addition to estimates of various IMs.

As stated by Douglas et al. (2015), although the characterization of earthquake shaking by a single number (an IM) is a great simplification, it makes seismic hazard assessment much more straightforward since the link between the seismic-source and ground-motion models can be expressed as a closed-form equation [ground motion prediction equations (GMPEs), also known as attenuation relation(ship)s] to estimate the probability of exceeding a given level of earthquake shaking. These probabilities are calculated through probabilistic seismic hazard assessment (PSHA) (Cornell, 1968; McGuire, 1976), which is the basis of most current seismic design maps, e.g. the National Annexes of Eurocode 8 (Comité Européen de Normalisation, 2005) and ASCE-7 (ASCE, 2013). Consequently it is still common to assess seismic hazard using PSHA through ground-motion models that return IMs. Then, based on this analysis and if needed, to obtain earthquake time-histories for the most important scenarios, generally defined using disaggregation (Bazzurro and Cornell, 1999), either through selection from a databank of natural accelerograms (NIST, 2011) or simulations of artificial records (Douglas and Aochi, 2008).

Because of the key role they still play in seismic hazard assessment, this review focuses on GMPEs derived empirically (i.e. from seismograms of real earthquakes). The purpose of this article is not to repeat the historical review of empirical ground motion estimation presented by Douglas (2003a) nor the overall scope of the review of all methods for ground-motion prediction by Douglas and Aochi (2008). Rather, this article seeks to review the great advances in ground-motion prediction over the past decade and to provide the reader with an overview of the principal topics of research. The

52 article concludes with some recommendations for future developments.

53 Although much of the following discussion concerns topics that are rel-  
54 evant for all tectonic regimes (e.g. shallow active crustal, subduction and  
55 stable continental) the examples are mainly taken from studies related to  
56 ground motions in shallow active crustal environments. A review focused  
57 on other tectonic regimes may emphasize other issues (e.g. the importance  
58 of focal depth for subduction events and simulation-based ground-motion  
59 models for stable continental regions). The wealth of data from shallow  
60 active crustal areas means that epistemic uncertainties are probably lower  
61 than in other tectonic regimes (e.g. Douglas, 2010b, Compare Figures 2, 8  
62 and 10). For instance, in some tectonic regimes (e.g. oceanic crust, deep  
63 Vrancea-type and the Himalaya) there are few strong-motion observations to  
64 constrain ground-motion models and consequently the epistemic uncertainty  
65 for these regions is much higher than for shallow active crustal areas.

## 66 **2. Summary of current state of practice**

67 It has now been more than fifty years since the first ground-motion model  
68 accounting for both magnitude and distance dependence was derived (Es-  
69 teva and Rosenblueth, 1964). Models are currently published at the rate of  
70 more than one per month and, at the last count, the total number of empir-  
71 ical equations for the prediction of PGA was 400 with many more based on  
72 simulations (Douglas, 2016). The close match between the rate of increase  
73 in strong-motion recordings and the number of GMPEs is shown in Figure 1.  
74 The rapidly increasing number of GMPEs led Bommer et al. (2010) to rec-  
75 ommend criteria for the selection of GMPEs to retain only those models for  
76 consideration that could be thought of as representing the state of the art.

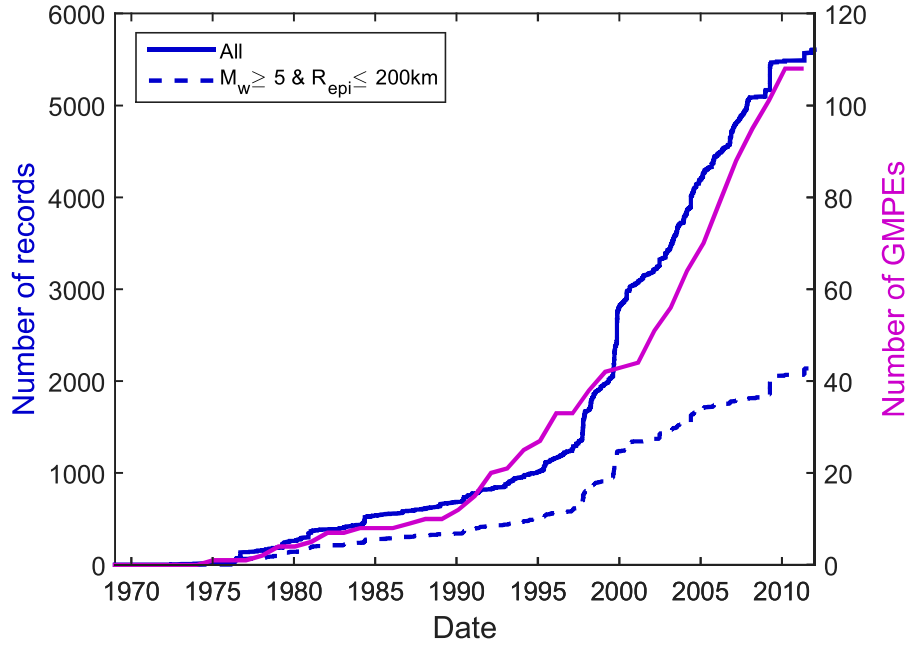


Figure 1: Available strong-motion records from RESORCE (Akkar et al., 2014b) (left-hand axis) and number of published GMPEs from Douglas (2016) (right-hand axis) against date for Europe and the Middle East (up to 2012).

They also suggest that these criteria could be used as a quality assurance step to guide publication of new GMPEs.

A brief comparison between the first ground-motion model (Esteva and Rosenblueth, 1964) and the recently-published GMPE of Abrahamson et al. (2014) helps demonstrate the developments in this field. The GMPE of Esteva and Rosenblueth (1964) was based on only 46 records and its three coefficients were estimated via standard least-squares regression. In contrast the model of Abrahamson et al. (2014) is based on over 15 000 records from more than 300 earthquakes and its roughly 40 coefficients were determined based on random-effects regression (Abrahamson and Youngs, 1992) or con-

87 strained based on ground-motion simulations or physical reasoning. Little  
88 information is provided on the data behind the model of Esteva and Rosen-  
89 blueth (1964) and it is thought that these data were obtained from various  
90 sources with seemingly little regard to their consistency or validity. In con-  
91 trast, the model of Abrahamson et al. (2014) is the outcome of careful data  
92 collection via the NGA projects (Power et al., 2008; Bozorgnia et al., 2014).  
93 The GMPE of Esteva and Rosenblueth (1964) is only for PGA and PGV  
94 because before the Caltech Blue Books (Brady et al., 1973) response spec-  
95 tra were difficult to obtain; whereas the model of Abrahamson et al. (2014)  
96 provides predictions for PGA, PGV and pseudo-SA at 22 periods between  
97 0.01 and 10 s. Finally, as is common for early GMPEs, Esteva and Rosen-  
98 blueth (1964) do not report the standard deviation ( $\sigma$ ) of their equation;  
99 whereas Abrahamson et al. (2014) concentrate much of their effort on de-  
100 riving a complex  $\sigma$  that models the different components of ground-motion  
101 variability.

102 In the decade or so since the review by Douglas (2003a) GMPE devel-  
103 opers have concentrated on: improvements in the estimation of the ground-  
104 motion variability associated with their models and its components (see  
105 Section 5); a move away from simple regression-based curve fitting; at-  
106 tempts at using non-parametric techniques; the use of much more and higher  
107 quality data; attempts at including additional independent parameters (see  
108 Section 3); a better appreciation of epistemic uncertainty (see Section 6);  
109 extensions of spectral models to shorter ( $< 0.1$  s) and longer ( $> 2$  s) peri-  
110 ods using individually-processed<sup>1</sup> records, often from digital instruments; a

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<sup>1</sup>The extension to shorter periods is aided by the observation (Douglas and Boore, 2011; Bommer et al., 2012) that SA is relatively unaffected by high-cut filtering.

more careful consideration of how the models perform beyond their ‘comfort zones’, e.g.: for  $\mathbf{M} < 5$ ,  $\mathbf{M} > 7$  and  $R < 10$  km; and making the models easier to use and test within PSHA (see Section 4). In addition, there has been a growing interest in developing models for other IMs, e.g. peak ground displacement, Arias intensity and various duration measures (see Section 7).

## 2.1. Current de facto standards

As demonstrated by the review of Douglas (2003a) many different choices, in terms of dependent and independent variables, derivation technique and functional form, were made by GMPE developers until the 1990s. In the past couple of decades, however, there has been a general convergence to a set of *de facto* standards.

Most developers now present models for PGA, increasingly PGV, and pseudo-SA for 5% of critical damping based on the geometric mean of the values from two horizontal components, or the orientation-independent horizontal component (Boore et al., 2006). They often use records from public online databases (e.g. Akkar et al., 2014b; Chiou et al., 2008) that have been low-cut filtered with record-specific cut-offs that are then respected when considering the reliable frequency ranges of their models.

The size of an earthquake is invariably characterized in terms of moment magnitude ( $\mathbf{M}$ ), although this is sometimes estimated from other magnitudes, commonly local magnitude ( $M_L$ ) (e.g. Bindi et al., 2005; Goertz-Allmann et al., 2011), duration magnitude ( $M_d$ ) (e.g. Bakun, 1984; Edwards and Douglas, 2014) or surface wave magnitude ( $M_s$ ) (e.g. Ambraseys and Free, 1997), through region-specific equations. Generally the earthquake is characterized into three faulting mechanisms (styles of faulting): normal, strike-slip and reverse. It is now common to consider nonlinear magnitude

137 scaling (see Section 4.2).

138     The length of the travel path from source to site is generally measured  
139 either in terms of the distance to the surface projection of the rupture (the so-  
140 called Joyner-Boore distance,  $r_{jb}$ ) (Joyner and Boore, 1981) or, accounting  
141 for the depth, the distance to the causative fault (the so called rupture dis-  
142 tance,  $r_{rup}$ ). For smaller earthquakes, where point sources can be assumed,  
143 these distance metrics become equal to epicentral ( $r_{epi}$ ) and hypocentral  
144 ( $r_{hyp}$ ) distances, respectively. Some recent studies present models for both  
145 finite-fault ( $r_{rup}$  or  $r_{jb}$ ) and point-source ( $r_{epi}$  or  $r_{hyp}$ ) distance metrics so  
146 that the correct GMPE is available when used within PSHA for point sources  
147 (e.g. within area sources) (Bommer and Akkar, 2012) without having to per-  
148 form conversions. It is also common to account for magnitude-dependent  
149 decay of IMs with distance (see Section 4.2).

150     Because boreholes were typically drilled to 30 m and because of its subse-  
151 quent use within many projects and design codes, e.g. Eurocode 8, the time-  
152 average shear-wave velocity in the top 30 m ( $V_{s,30}$ ) is the common way that  
153 near-surface site conditions are characterized within recent GMPEs, either  
154 directly or, when insufficient information is available, through site classes.  
155 It is still relatively uncommon for GMPEs to account directly for potential  
156 nonlinear site amplification because this behavior is rare within observed  
157 strong ground motions. Within PSHA non-linear effects generally require  
158 a simulation-based site term to be adopted, often from a stand-alone study  
159 (Kamai et al., 2014; Seyhan and Stewart, 2014; Sandikkaya et al., 2013).

160     Finally it has become standard to use either random-effects (Abraham-  
161 son and Youngs, 1992) or one- or two-stage maximum-likelihood regression  
162 (Joyner and Boore, 1993) to estimate the free coefficients of the model.  
163 These techniques, applied to the same data, would lead to very similar

164 results, although the latter may be more susceptible to trade-offs. Both  
165 techniques provide estimates of the between- and within-event components  
166 of ground-motion variability (see Section 5).

### 167 **3. Additional independent variables**

168 To obtain GMPEs that estimate more appropriate ground motions for a  
169 given earthquake, path and site, independent variables in addition to mag-  
170 nitude, faulting mechanism, source-to-site distance and a near-surface site  
171 class (or  $V_{s,30}$ ) have been tested and/or included within some recent models.  
172 These attempts are briefly discussed in this section.

#### 173 *3.1. Source parameters*

174 All GMPEs include magnitude as the main source parameter. This is  
175 now routinely moment magnitude due to its robustness, the fact that it  
176 does not saturate, and because it is possible to estimate from historical and  
177 palaeological information. The latter consideration is important in linking  
178 GMPEs to earthquake catalogs, where the longer the available time-period  
179 the more reliable are recurrence relations, particularly at higher magnitudes.  
180 While magnitude is certainly an important factor for ground-motion ampli-  
181 tudes, there are other source parameters that can control the amplitude and  
182 frequency content of radiated seismic energy. The most influential of these  
183 is the earthquake stress drop. While the stress drop has a physical mean-  
184 ing, there are different definitions (e.g. static, dynamic or ‘Brune’). When  
185 referred to in engineering seismology applications ‘stress drop’ or ‘stress  
186 parameter’ is effectively used to refer to the proportion of high-frequency  
187 energy (for a given magnitude) that is radiated from the source (Atkinson  
188 and Beresnev, 1997).



189       Following on from observations of Somerville (2003), model developers of  
 190       the NGA West 1 and 2 projects (Power et al., 2008; Bozorgnia et al., 2014)  
 191       investigated the impact of depth to the top of the rupture plane ( $Z_{TOR}$ ) on  
 192       ground motions. Some of them (e.g. Campbell and Bozorgnia, 2014) find  
 193       that using  $Z_{TOR}$  within the model leads to statistically better predictions  
 194       with deep earthquakes generating higher ground motions than shallow events  
 195       (all other things being equal), which could be explained by higher stress  
 196       drops. Possible lower stress drops for aftershocks is behind the decision of  
 197       some NGA West developers to exclude data from this type of event (e.g.  
 198       Boore and Atkinson, 2008) whereas others (e.g. Chiou and Youngs, 2008)  
 199       include terms to account for this difference. This effect appears to be small  
 200       and could be related to the way that earthquakes are classified (Douglas and  
 201       Halldórsson, 2010). Radiguet et al. (2009) present evidence that SAs from  
 202       immature faults are statistically-significantly higher than those from mature  
 203       faults, which again could be related to higher stress drops for earthquakes  
 204       occurring on immature faults. The maturity of faults has yet to be included  
 205       in a GMPE because the age of faults is not a readily-available parameter.  
 206       The recent ground-motion model by Bora et al. (2015) includes an explicit  
 207       term for the stress (drop) parameter ( $\Delta\sigma$ ) commonly used within stochastic  
 208       models (e.g. Atkinson and Silva, 2000; Rietbrock et al., 2013), while Douglas  
 209       et al. (2013) and Bommer et al. (2016) present unique GMPEs for a range of  
 210        $\Delta\sigma$ . This allows models to be readily employed in areas where the average  
 211       stress drop is known but it puts the onus on the user to select an appropriate  
 212       median  $\Delta\sigma$  (and uncertainty about this value).

213       Directivity of earthquake ground motion fields is an emerging topic that  
 214       has been addressed, for example, in the recent NGA West 2 project (Spudich  
 215       et al., 2014). While often clear in large-magnitude earthquake simulations,

216 this issue has seen relatively little focus in recent years. This is primarily due  
217 to the nature of PSHA, which combines all possible earthquake scenarios:  
218 rupture directivity effects, therefore, tend to be smoothed out. However, in  
219 understanding deterministic hazard, or for future analyses, where rupture  
220 directivity preference can be assigned, accounting for this effect may help to  
221 reduce epistemic uncertainty.

### 222 3.2. *Path parameters*

223 Path terms within GMPEs have grown more complex in terms of their  
224 functional form over the past decade with the realization that ground mo-  
225 tions from small and large earthquakes do not decay at the same rate (see  
226 Section 4.2). In addition, because of the availability of ground-motion data  
227 (often from broadband instruments or high-sensitivity strong-motion sen-  
228 sors) at distances greater than 100 km (roughly the limit of analogue strong-  
229 motion recording) a number of GMPEs include terms to model anelastic  
230 attenuation, the rate of which is sometimes considered regionally-dependent  
231 (see Section 4). Cousins et al. (1999), for example, developed a GMPE  
232 for New Zealand that accounts for additional attenuation for travel paths  
233 through volcanic regions by including a term that is a function of the hori-  
234 zontal distance through such zones.

235 Nevertheless, commonly travel path is simply parameterized using source-  
236 to-site distance. This means that the decay rate is the same for all locations  
237 irrespective of the crustal structure. Douglas et al. (2004, 2007) develop a  
238 technique based on simulations to calculate an equivalent hypocentral dis-  
239 tance that captures the impact of crustal structure on ground-motion decay  
240 and, consequently, allows a ground-motion model to be branched into region-  
241 specific models. This approach has yet to be applied for the derivation of a

242 GMPE for use in practice.

243 A handful of GMPEs (generally for use in California) include terms to  
244 model the location of a site with respect to the hanging and foot walls of  
245 the causative fault (e.g. Campbell and Bozorgnia, 2014; Abrahamson et al.,  
246 2014), sometimes by using  $R_x$  (the horizontal, strike-normal distance to the  
247 shallowest part of the surface projection of the fault). The terms to model  
248 this effect are often complex and hence rely on simulations to constrain their  
249 free parameters. For applications in areas without clearly-defined dipping  
250 faults such terms are often turned off when the model is used within PSHA.

### 251 3.3. Site parameters

252 As discussed in Section 2.1, most current GMPEs use  $V_{s,30}$  or site classes  
253 based on  $V_{s,30}$  to characterize the near-surface conditions at a site. In an  
254 attempt to account for the effect of deeper structure on ground motions,  
255 some recent GMPEs for California often use, in addition to  $V_{s,30}$ , either the  
256 depth to the 1 km/s velocity horizon ( $Z_{1.0}$ ) (e.g. Chiou and Youngs, 2014)  
257 or the depth to the 2.5 km/s horizon ( $Z_{2.5}$ ) (e.g. Campbell and Bozorgnia,  
258 2014).  $Z_{1.0}$  and  $Z_{2.5}$  are often strongly correlated but weakly correlated  
259 with  $V_{s,30}$  and hence their use alongside  $V_{s,30}$  adds discriminatory power to a  
260 GMPE. For many parts of the world estimates of  $Z_{1.0}$  and, particularly,  $Z_{2.5}$   
261 are, however, difficult to obtain because they require knowing the shear-wave  
262 velocity down to hundreds or thousands of meters. Consequently, empirical  
263 relationships to estimate these parameters from  $V_{s,30}$  have been proposed  
264 (Boore et al., 2011) to center the predictions at an average  $Z_{1.0}$  or  $Z_{2.5}$ .

265 PSHA is often conducted for a rock site with  $V_{s,30}$  equal or larger than  
266 760 m/s [the NEHRP B/C boundary (National Earthquake Hazard Reduc-  
267 tion Program, 1994)] (see Section 4.4). At high  $V_{s,30}$  the site amplification

modeled in the GMPE will be low and any nonlinearity in modeled response  
 weak. One of the largest changes in PSHA for such sites in the past decade  
 has been the appreciation that site amplification related to shear-wave ve-  
 locity is not the whole story but that high-frequency attenuation, generally  
 modeled by  $\kappa$  (Anderson and Hough, 1984), also needs to be considered.  
 The effect of an average  $\kappa$  is implicitly captured within empirical GMPEs  
 through the data that are used. The average  $\kappa$  implied by the shape of the  
 short-period spectra of GMPEs evaluated for high  $V_{s,30}$  is, however, often  
 much higher than the  $\kappa$  measured at rock sites. Consequently, as discussed  
 in Section 4.5, a host-to-target adjustment for  $\kappa$  is required when these  
 GMPEs are used in a site-specific study. In an attempt to overcome this  
 requirement, Laurendeau et al. (2013) introduce a term for  $\kappa$  directly into  
 a GMPE developed from Japanese data. Use of such a model means that  
 $\kappa$  needs to be known for a site of interest. This is the apparent drawback  
 of introducing new variables into GMPEs: the requirement for the user to  
 know their value and their uncertainty for their study. In the past, however,  
 the user generally assumed that the implicit average value within the data  
 used to derive the GMPE was appropriate for their site.

#### 4. Regional models

With the rapidly-growing quantity of data from digital strong-motion  
 networks, which accurately record earthquakes down to  $M3$  and below, there  
 has been a move towards the development of GMPEs for small geographical  
 regions (e.g. national or sub-national) and partially away from models cov-  
 ering large tectonic regimes, e.g. shallow crustal earthquakes globally. An  
 idea of the utility of this approach for the development of empirical GMPEs

Table 1: The number of years required to record fifty  $M_w \geq 5$  shallow earthquakes assuming dense strong-motion network covering whole territory (country or state) based on the International Seismological Centre’s earthquake catalog from 1992 to 2012.

Country	Number of years
Japan	7
Turkey	9
Greece	12
California	20
Italy	31
Iceland	140
Spain	250
France	1000
United Kingdom	$\gg 1000$

given only data from a country or state can be gained from Table 1. For some highly seismically active areas this goal of purely-national GMPEs is feasible but for less active (e.g. Spain) or smaller countries (e.g. Iceland) local records would have to be used in conjunction with simulations or foreign data to derive robust models.

As discussed in Section 4.2, there are difficulties in developing regional models for use within standard seismic hazard assessments unless the models are derived using data from large events. Therefore, to account for potential regional dependency some GMPE developers derive a robust model using data from a variety of regions within a single tectonic regime (e.g. shallow crustal) and then add terms when required to account for observed regional differences. For example, Boore et al. (2014) include terms to model differences in anelastic attenuation in China/Turkey and Japan/Italy to other ar-

306 eas (predominantly California). In addition to regional variations in median  
307 predictions, the variability of ground motion may be regionally-dependent.  
308 For example, Abrahamson et al. (2014) differentiate between variability in  
309 Japan and elsewhere.

310 Regional dependence of ground-motion models is, therefore, still a topic  
311 of ongoing research. The issue is somewhat complicated by the sweeping  
312 terms typically used to classify tectonic regions: stable continental, shallow  
313 active crustal and so forth. Within each of these groups significant variability  
314 in both structure and geology exists – meaning that systematic variability  
315 in ground motion may be obscured if only looking at differences within or  
316 between these classes. Nevertheless, it is generally acknowledged that at dis-  
317 tances larger than around 50 km, regional variations in geology and tectonic  
318 structure lead to significant differences in ground motion attenuation (e.g.  
319 Boore et al., 2013; Kotha et al., 2016b,a). On the other hand, differences  
320 at shorter distances are less well understood due to limited data and the  
321 complexity of earthquake sources. Regional differences in stress fields due  
322 to factors such as tectonic loading and structure (Gölke and Coblentz, 1996),  
323 or, at smaller scales, due to fault structure and maturity (Manighetti et al.,  
324 2007) may lead to differences in earthquake stress drop that can be observed  
325 at national (e.g. Goertz-Allmann and Edwards, 2014) or local scales (e.g.  
326 Allmann and Shearer, 2007). The resolution of such analyses is, however,  
327 debated due to the trade-off with attenuation, which is typically assumed to  
328 be homogeneous. Addressing the issue of regionalization of ground-motion  
329 models requires more data, particularly at short distances. In the meantime,  
330 hazard analysts can use hazard disaggregation to understand, to a first or-  
331 der, the sensitivity of possible regional ground motions on seismic hazard.  
332 For instance, hazard is often primarily driven by relatively close earthquakes

333 ( $< 50$  km) and, hence, regional differences in geology will be less important  
334 to understand than differences in fault-rupture kinematics, for example.

#### 335 4.1. *Testing of GMPEs*

336 When conducting a seismic hazard assessment for a region that is not  
337 covered by a selected GMPE it has been increasing common to undertake  
338 a quantitative comparison between predictions and the ground motions ob-  
339 served in the region (Stewart et al., 2015). This has only become possible  
340 for many parts of the world since the advent of digital ground-motion net-  
341 works in the past couple of decades. Various methods have been developed  
342 to undertake this testing but they are invariably based on ‘residuals’<sup>2</sup>, either  
343 total or, more correctly, separated into between- and within-event compo-  
344 nents (Stafford et al., 2008), between predictions and observations. The  
345 most employed techniques are those by Scherbaum et al. (2004), Scherbaum  
346 et al. (2009) and Kale and Akkar (2013). A more informative approach is to  
347 consider plots of the residuals with respect to magnitude, distance and other  
348 variables to understand what parts of the model are causing any misfits (e.g.  
349 Scasserra et al., 2009).

350 A difficulty with such testing is that it is difficult to judge how much  
351 weight should be given to a good or poor match as the available data are  
352 often sparse and/or only available for magnitude and distance ranges of  
353 limited engineering interest (Beauval et al., 2012). If a poor match is found  
354 between observations and predictions and this is judged to be robust then  
355 adjustment factors can potentially be derived to modify the GMPE so that

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<sup>2</sup>They are not strictly residuals because generally the data compared were not used for the derivation of the tested GMPE.

356 it provides better predictions (Bommer et al., 2006). This approach has  
357 been formalized in the so called referenced-empirical technique by Atkinson  
358 (2010) and variants of it have been applied in various projects, particularly  
359 to adjust models for small and moderate events (e.g. Bourne et al., 2015).

#### 360 *4.2. Scaling of ground motions for small and large earthquakes*

361 In the past decade there has been a push to derive GMPEs to predict  
362 accurately ground motions from earthquakes with  $M < 5$ . Until the estab-  
363 lishment of digital strong-motion networks, which started in many regions  
364 in the late 1990s, ground-motion databases generally became sparse below  
365 about  $M5$ . In addition, for high seismicity areas, where most of the available  
366 data are from, the dominant earthquake scenarios for engineering purposes  
367 are generally at  $M > 5.5$ . Consequently there was little call for GMPEs  
368 that could be used confidently for small earthquakes.

369 The development of such models in the past decade has been driven  
370 by the availability of large sets of records from digital networks with good  
371 coverage down to often  $M3$  for many parts of Europe and elsewhere. Often  
372 these data are used to derive regional GMPEs (see Section 4) generally  
373 without the inclusion of data from larger earthquakes. When applying a  
374 GMPE in a different geographical region than for which it was originally  
375 derived it is important to check it against local data. As shown by, for  
376 example, Douglas (2003b), unless the GMPE was derived using data from  
377 small events and an appropriate functional form was used there will likely  
378 be a large discrepancy between predictions and observations. This has been  
379 used as an argument for a strong regional dependency in ground motions  
380 but, as shown by Cotton et al. (2008) amongst others, it is likely due to the  
381 differing magnitude ranges of the observations and model. Another recent



382 driver in the development of GMPEs that cover the range below **M5**, even  
383 for high seismicity zones, is the need for such models to estimate components  
384 of the ground-motion variability that require many records from the same  
385 site (see Section 5.3).

386 As shown by Douglas (2003b, Figure 4), Douglas and Jousset (2011)  
387 and Baltay and Hanks (2014), empirical GMPEs derived from data from  
388 small earthquakes generally show higher dependency on magnitude, partic-  
389 ularly for short-period IMs, than those models derived for moderate and  
390 large events. This means that extrapolation of these models beyond the  
391 magnitude range for which they were derived often leads to over-prediction.  
392 Fukushima (1996), Douglas and Jousset (2011) and Baltay and Hanks (2014)  
393 demonstrate that a simple stochastic model (Boore, 2003) with a single-  
394 corner source spectrum (Brune, 1970) and high-frequency attenuation (An-  
395 derson and Hough, 1984) reproduces the observed magnitude-scaling of em-  
396 pirical GMPEs and demonstrates why extrapolation of such models is so  
397 problematic. Algorithmic differentiation (Molkenthin et al., 2014) can be  
398 used to study the scaling of GMPEs with respect to its input parameters,  
399 which aids understanding of how the models behave and extrapolate.

400 As well as magnitude-scaling being different for ground motions from  
401 small and large earthquakes, the decay with distance also differs. Earth-  
402 quake magnitude has two effects on the distance dependence of ground-  
403 motion attenuation. The first is due to near-field saturation: as one ap-  
404 proaches a finite source, the contribution from the far ends of the source  
405 become increasingly small due to the distance that the energy must propa-  
406 gate to reach you (attenuation effects) and the time which this takes (scat-  
407 tering and dispersion effects). At short and moderate structural periods,  
408 therefore, the peak amplitudes of a **M7** event are similar to an **M8**. The

primary difference is the duration and spatial extent over which the motions occur, being significantly longer and more widespread in the latter case. The second effect is the distance dependence of the ground motion decay. For increasingly large events the finite nature of the source means that ground motion does not decay as quickly as for small (roughly point) sources, since the motion at distance is increased by constructive interference from later arrivals along the finite fault (e.g. Boore, 2009). In fact, even for point-source models, Cotton et al. (2008) showed that the decay of response spectral ordinates is magnitude-dependent due to the influence of spectral shape. To capture this, functional forms of GMPEs in the past decade have often used magnitude-dependent decay terms.

#### 4.3. *Non-tectonic earthquakes*

Although the vast majority of GMPEs are still derived for tectonic earthquakes, a growing number of models are available for earthquakes of other types, e.g. those induced by mining (e.g. McGarr and Fletcher, 2005) or fluid injection (e.g. Douglas et al., 2013). Seismic hazard assessments for human-activity-related, induced or triggered earthquakes require ground-motion models that are adapted to this type of event and it is not *a priori* clear that shaking from such shocks is similar to that from natural earthquakes. In addition, the magnitude, source-to-site distance and focal depth range of importance for induced seismicity is generally smaller than the focus of hazard assessments for natural earthquakes. Hence, as discussed in Section 4.2, this leads to the need to develop models to account for this difference. The finding of Douglas et al. (2013) that motions from induced and natural shallow seismicity are statistically similar means that the more abundant data banks of records from small natural shallow earthquakes could be

435 used to derive GMPEs for use within hazard assessments for induced seismic-  
436 ity (e.g. Atkinson, 2015). It could also be argued that with an appropriate  
437 correction for depth [i.e. for distance and stress-drop (Hough, 2014)], data  
438 from deeper natural seismicity could be used to determine ground-motion  
439 fields of larger induced events.

#### 440 *4.4. Prediction for a reference velocity horizon*

441 Ground motion within PSHA is typically estimated for a reference site,  
442 circumventing the geological heterogeneity of the uppermost layers. This  
443 is often at or around the NEHRP class B/C boundary of 760 m/s or the  
444 Eurocode 8 class A/B boundary of 800 m/s (e.g. Delavaud et al., 2012).  
445 Subsequently, the results of microzonation or site-specific response analyses  
446 can be applied in conjunction with these estimates. The reason for this is the  
447 significant variability of resolution, reliability and availability of site-specific  
448 data. Practitioners are, in this way, free to apply their own site specific  
449 corrections to a regionally-consistent hazard map for reference rock.

450 Site response terms within GMPEs are included for two reasons. Firstly,  
451 to enable ground-motion records from all site conditions (including non-  
452 rock stations, which comprise the majority of most strong-motion networks)  
453 to be used to derive GMPE that would be statistically more robust than  
454 using only rock records. A few developers (e.g. Idriss, 2014) exclude records  
455 from sites with low  $V_{s,30}$  because they believe that it is not possible to  
456 capture site response by means of a simple site term. Consequently such  
457 models are generally based on far fewer records but the risk of bias from  
458 site amplification is reduced. The second reason for including site terms in  
459 GMPEs is that such models allow seismic hazard assessments for a variety  
460 of sites (including non-rock sites) to be easily conducted, which could be

461 useful when high accuracy is not a requirement.

462 In a similar way, recent PSHAs (e.g. Bommer et al., 2015) predict the  
463 ground motion initially at a subsurface reference rock horizon, choosing a  
464 depth below which lateral variability is considered insignificant (usually at a  
465 wave velocity consistent with ‘engineering’ or hard rock). Site-specific non-  
466 linear amplification is then applied during the hazard calculation based on  
467 site-response analyses. This approach has the benefit of potentially reducing  
468 the site-to-site variability in predicted ground motion. If one assumes the  
469 full range of site variability is captured through this process then the GMPE  
470 component of site-to-site variability  $\phi_{S2S}$  (see Section 5.3) can be set to zero,  
471 leading to non-ergodic single-station sigma (Atkinson, 2006). Practitioners  
472 must be careful in this case that the modeled variability of the site response  
473 is sufficient, but at the same time not so high that ergodic  $\sigma$ s are exceeded  
474 due to uncertainty in site response analyses.

475 The move towards reference-site hazard and reference horizons to make  
476 best use of site-response analyses means that GMPEs are being increasingly  
477 evaluated for relatively high  $V_{s,30}$  (e.g.  $\geq 760$  m/s). This is one of the factors  
478 driving the derivation of new GMPEs. Sites with high  $V_{s,30}$ , however, are  
479 poorly represented in strong-motion databases because many stations are  
480 installed in urban environments on soft and stiff soils (e.g. Akkar et al.,  
481 2010).

#### 482 4.5. Host-to-target adjustments

483 Ground motion is dependent on the shear-wave velocity and attenuation  
484 characteristics of the upper layers of soil and rock. When modifying site  
485 conditions, e.g. changing predictions relevant for California to a site-specific  
486 target in the United Kingdom, hazard analysts must consider the effect of

487 this change on the predicted ground motion. This is done through host-to-  
488 target adjustments.

489 As stated above, GMPEs are typically developed using site descriptors  
490 such as class (e.g. rock, stiff soil and soft soil) or  $V_{s,30}$ . It is important  
491 to note, however, that when using a GMPE estimates are implicitly tied  
492 to a range of possible site types that fall within the site descriptor and  
493 this may be biased by a particular geology. Even GMPEs using  $V_{s,30}$  will  
494 cover a range of site types because many velocity profiles are possible for a  
495 given  $V_{s,30}$ . While different velocity profiles can lead to the same  $V_{s,30}$ , they  
496 may lead to significantly different amplifications (e.g. Castellaro et al., 2008;  
497 Papaspiliou et al., 2012). If a particular velocity structure (e.g. low velocity  
498 soils over a high velocity basement) is characteristic of a region, then ground  
499 motion at a  $V_{s,30}$  in one region may be systematically different to that in  
500 another with a different average structure. As discussed previously, some of  
501 this site variability can be captured by using additional site parameters, such  
502 as  $Z_{1.0}$  or  $Z_{2.5}$ . Recent PSHA studies have, however, moved towards fully  
503 accounting for the effect of site-specific characteristics, by taking advantage  
504 of the wealth of information often available for site-specific hazard analyses.  
505 Such differences are accounted for by using host-to-target adjustments. The  
506 same approach can be used to modify ground-motion predictions made at a  
507 particular  $V_{s,30}$  and provide them at another. This approach is particularly  
508 useful in the case that GMPE predictions are considered unreliable at the  
509 target  $V_{s,30}$ .

510 Since earthquake engineering generally uses SA, direct adjustments of the  
511 Fourier amplitude spectra (FAS) cannot be used to perform host-to-target  
512 adjustments. This is because ground motion at a given oscillator period is  
513 dependent not only on the FAS at that period but also other values around

514 it (e.g. Bora et al., 2015). The host-to-target ratio is, therefore, dependent  
 515 on the input ground motion in addition to the different site properties. The  
 516 hybrid-empirical method (HEM) based on Campbell (2003) is commonly  
 517 used to make host-to-target adjustments. HEM uses stochastic simulations  
 518 [typically using random-vibration theory (RVT) (Cartwright and Longuet-  
 519 Higgins, 1956)] to generate FAS-compatible response spectra for the host  
 520 and target sites, which can then be used to calculate the ratio in terms of  
 521 SA.

522 Using RVT through the HEM allows transformations from the Fourier  
 523 domain into the response spectral domain. HEM, however, requires a full  
 524 seismological model (for source, path and site) of the host and target re-  
 525 gions. Because of this Al Atik et al. (2013) developed a method based on  
 526 inverse RVT (IRVT) (Vanmarcke and Gasparini, 1976) that can be used to  
 527 modify response spectra for host-to-target adjustments in the Fourier do-  
 528 main. The method has the advantage that no assumptions on the form  
 529 of the host model (GMPE) are required. Working in the Fourier domain  
 530 has the advantage that adjustments are independent of the input motion  
 531 unlike when working in the response spectral domain. For a given signal  
 532 duration (often defined based on simple regional models), IRVT transforms  
 533 the response spectrum into a compatible FAS. FAS based host-to-target  
 534 conversion can then be applied to the response-spectrum-compatible FAS  
 535 before being returned to the response domain through the standard RVT  
 536 approach. A limitation of the IRVT approach is that the response spectrum  
 537 becomes less sensitive to the FAS as oscillator period decreases. This results  
 538 in significant non-uniqueness of the response-spectrum-compatible FAS at  
 539 short periods (roughly  $T < 0.05$ s). Nevertheless, an advantage of this ap-  
 540 proach is that one can directly estimate seismological parameters from the

541 GMPE-compatible FAS, such as  $\kappa$ .

542 Figure 2 shows an application of the  $V_s$ - $\kappa_0$  corrections to GMPEs used  
543 in the Swiss National Seismic Hazard Maps (Edwards et al., 2016). The  
544 selected target  $V_s$  profile (Poggi et al., 2011,  $V_{s,30} = 1105$  m/s) and  $\kappa_0$  value  
545 (Edwards et al., 2011,  $\kappa_0 = 0.016$  s) define the reference rock for the seismic  
546 hazard map. For each GMPE two possible host  $V_s$  profiles were selected  
547 (with defined  $V_{s,30}$  where the GMPE’s developers considered the best data  
548 coverage for rock). Four  $\kappa_0$  values were also selected for each GMPE using  
549 either  $V_{s,30}$ - $\kappa_0$  correlations or direct measurement using IRVT. The resulting  
550 eight  $V_s$ - $\kappa_0$  corrections for each GMPE were considered to represent the  
551 epistemic uncertainty involved in adjusting GMPEs to the regional reference.  
552 Small but significant differences arise at long periods due to differences in  
553 amplification of the host- $V_s$  profiles. Far more significant, however, is the  
554 epistemic uncertainty evident in the correction at short periods ( $T < 0.1$  s),  
555 which is due to the uncertainty in defining  $\kappa_0$  (e.g. Edwards et al., 2015).  
556 Similar observations are made by Rodriguez-Marek et al. (2014) for a site-  
557 specific hazard assessment.

## 558 5. Aleatory variability

559 Over the past decades there has been a growing realization that predict-  
560 ing shaking in future earthquakes is associated with large uncertainties and  
561 that this uncertainty must be captured within seismic hazard assessments.  
562 It has become standard to split these uncertainties into two components:  
563 those of inherent randomness, referred to as aleatory variability (this sec-  
564 tion) and those relating to a lack of knowledge or understanding, referred  
565 to as epistemic uncertainty (Section 6).

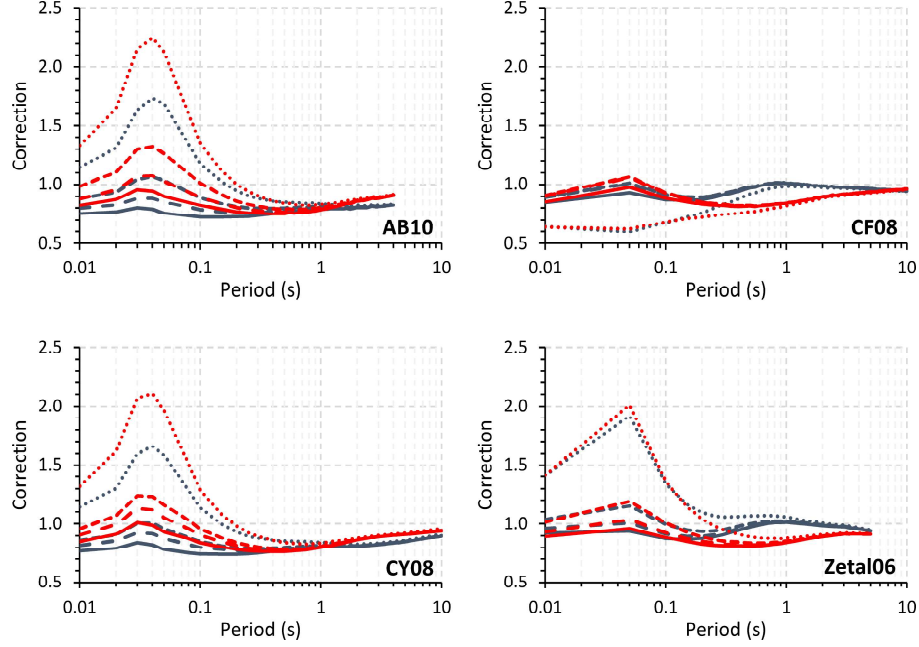


Figure 2:  $V_s$ - $\kappa_0$  corrections proposed for the Swiss National Seismic Hazard Maps by Edwards et al. (2016). Blue/Red indicate different host  $V_s$  profiles (two for each GMPE), line types indicate different  $\kappa_0$  (four for each GMPE) resulting in eight possible corrections per GMPE. AB10: Akkar and Bommer (2010); CF08: Cauzzi and Faccioli (2008); CY08: Chiou and Youngs (2008); and Zetal06: Zhao et al. (2006). The target properties are  $V_{s,30} = 1105$  m/s and  $\kappa_0 = 0.016$  s.



566 The definition of aleatory (and consequently epistemic) variability in-  
 567 evitably leads to disagreement and confusion. It could be argued, for in-  
 568 stance, that given a perfect model, aleatory variability is, by definition,  
 569 zero. However, in current understanding we can at least separate the vari-  
 570 ability into parts that can be quantified in terms of scientific uncertainty (e.g.  
 571 using different models to predict the same phenomena, such as site ampli-  
 572 fication), and those for which there is (at least currently) no scientifically-  
 573 based predictive capability (e.g. the stress-drop of the next earthquake). A  
 574 more appropriate terminology may therefore be *apparent* aleatory variabil-  
 575 ity with respect to a chosen model (written communication, J. J. Bommer,  
 576 2016). The advantage of splitting uncertainty into constituent components  
 577 is that the logic-tree approach (Kulkarni et al., 1984) can then be used  
 578 to branch through the epistemic uncertainty space (e.g. by selecting and  
 579 weighting different models) and allowing site or region-specific selections to  
 580 be made along with sensitivity studies and analyses (e.g. disaggregation) at  
 581 a branch-by-branch level. The distinction between aleatory and epistemic  
 582 is particularly important, for example, in the case of a fully probabilistic  
 583 seismic risk (or safety) assessment for a safety critical structure such as a  
 584 nuclear power plant. Such assessment requires the fractiles of the hazard  
 585 to be defined, which can only be correctly calculated with an appropriate  
 586 separation of aleatory and epistemic uncertainty.

587 Following Douglas (2003a), Strasser et al. (2009) observe that  $\sigma$  associ-  
 588 ated with GMPEs has shown little or no decrease since the 1970s despite  
 589 the increasing complexity of models. This fact and the importance of  $\sigma$  on  
 590 the results of PSHAs at long return periods, has encouraged attempts to  
 591 increase the complexity of models to account for other effects than simply  
 592 magnitude, distance and site class (see Section 3). To date these attempts

have not led to significant reductions in  $\sigma$  because GMPEs remain simple representations of complex physical phenomena. Improvements to metadata do, however, lead to slight reductions in assessed  $\sigma$ . For example, the model of Chiou and Youngs (2014) is associated with a smaller  $\sigma$  when measured  $V_{s,30}$  is used for a site than when an estimate of this site parameter is employed.

One of the major areas of engineering seismology research in the past decade has been in separating  $\sigma$  into its different components (Al Atik et al., 2010; Lin et al., 2011; Rodriguez-Marek et al., 2013) and using the appropriate components when conducting a hazard assessment (e.g. Walling and Abrahamson, 2012). There has also been a move from using whatever data were available towards selecting to: limit bias, exclude unreliable data, make analysis easier, and obtain more reliable  $\sigma$  estimates. As noted above, it has become standard to use random-effects/maximum-likelihood methods to estimate between-event ( $\tau$ ) and within-event ( $\phi$ ) components.

Records from nearby sites are correlated, which has been recognized by Jayaram and Baker (2010) when developing a regression technique to account for spatial correlations and by Boore et al. (1993), who choose only a single record per site class within a radius of 1 km. These spatial correlations are also important when conducting PSHA for infrastructure with considerable spatial extent or when computing group earthquake risk over an extended area.

### 5.1. *Between-event variability*

Aleatory variability within a given GMPE is usually separated into between- and within-event components ( $\tau$  and  $\phi$ , respectively). Between-event terms (random-effects in the context of random-effects regressions),

619 which are source-specific, are thought to be mainly related to stress drop  
 620 (Cotton et al., 2013). Using stochastic simulations, Drouet and Cotton  
 621 (2015) showed that the between-event variability was strongly controlled  
 622 by the stress parameter (as noted previously, ‘stress parameter’ is used to  
 623 avoid physical interpretation in terms of pure ‘stress drop’ and rather in-  
 624 dicate the proportion of high-frequency energy radiated by an earthquake).  
 625 The between-event term can, therefore, be thought of as describing how  
 626 energetic the rupture was compared to the average for a given magnitude  
 627 (all other things being equal). Such features are not possible (currently)  
 628 to predict and, therefore, fall into the category of aleatory variability. The  
 629 standard deviation of these event terms is described by  $\tau$ .

630 One of the main ways GMPEs are improving is related to the record-  
 631 ing of each earthquake by an increasing number of stations (in particular,  
 632 fewer singly-recorded events) so that the source terms (and  $\tau$ ) are better  
 633 constrained. This is particularly true for models based on predominantly  
 634 Californian or Japanese data but much less so for models derived from data  
 635 from Europe and the Middle East (Table 2 and Figure 3). This shows  
 636 that despite recent improvements in strong-motion networks in Europe and  
 637 Middle East, strong motion databases there remain dominated by poorly-  
 638 recorded events. For models based on data with low record-to-event ratios  
 639 the source terms (e.g. style-of-faulting factors) and  $\tau$  are poorly constrained.  
 640 Additionally, the small number of well-recorded events have a strong influ-  
 641 ence on the model.

642  $\tau$  is often found to be heteroscedastic, with decreasing variability as mag-  
 643 nitude increases (e.g. Youngs et al., 1995) (Figure 4). Estimated ground-  
 644 motion variability from small events ( $M < 5$ ) is often significantly larger  
 645 than at moderate and large magnitudes, with many GMPE developers avoid-

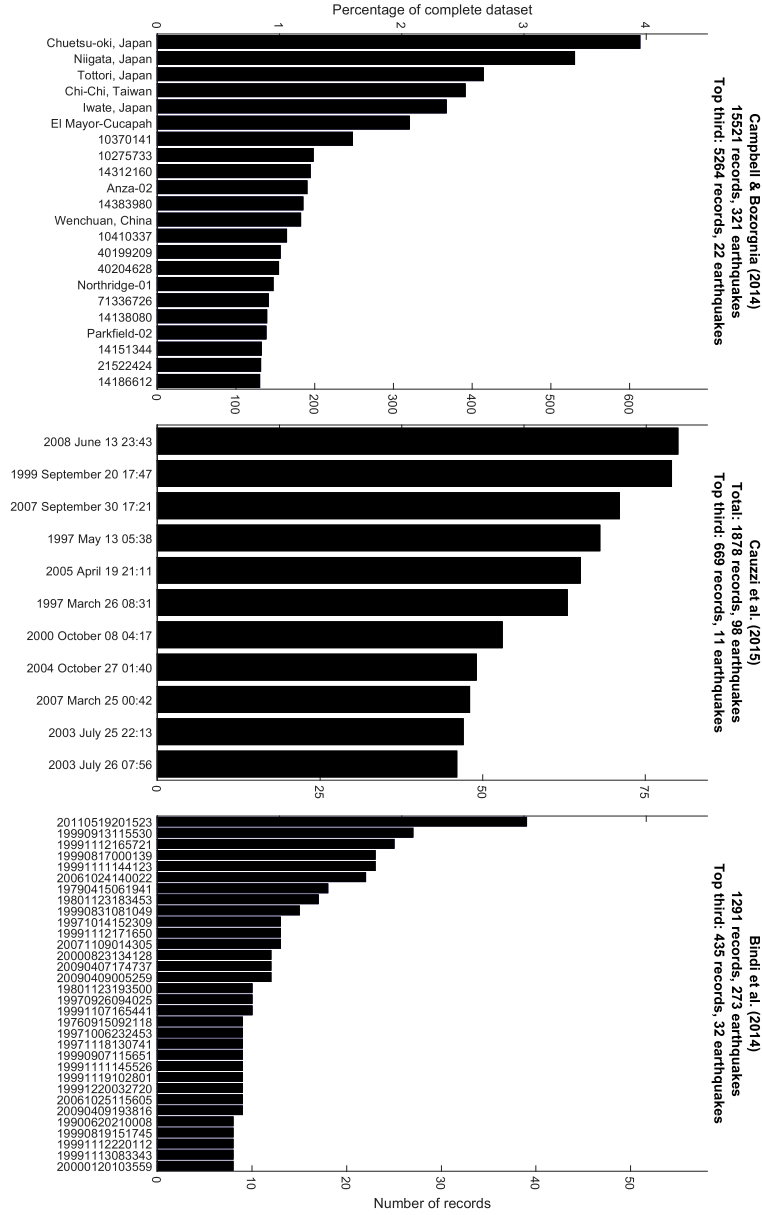


Figure 3: Number of records (bottom axes, different scales for all three subplots) and percentage of total (top axes, same scales for all three subplots) from earthquakes contributing to the top third of total number of records to three recent GMPEs: Campbell and Bozorgnia (2014) (predominantly Californian data), Cauzzi et al. (2015) (predominantly Japanese data) and Bindi et al. (2014) (European and the Middle Eastern data).

Table 2: Ratio (R/E) of number of records (R) per event (E) for four generations of ‘Californian’ and ‘European’ models.

‘Californian’ model	R	E	R/E	‘European’ model	R	E	R/E
Joyner and Boore (1981)	182	23	8	Ambraseys and Bommer (1991)	529	219	2
Boore et al. (1997)	271	20	14	Ambraseys et al. (1996)	422	157	3
Boore and Atkinson (2008)	1574	58	27	Ambraseys et al. (2005)	595	135	4
Boore et al. (2013)	~15000	~350	43	Akkar et al. (2014a)	1041	221	5

ing using data from small earthquakes. This is despite the need for models at lower magnitudes, e.g. for seismic hazard assessment from induced seismicity, to examine the applicability of a GMPE in a new region and to study the various components of ground-motion variability. While models of ground-motion variability have improved significantly in recent years, we must be careful not to over-interpret features of these models due to the limitations of separating the different contributions. In Figure 4 there is a peak at 0.1 s for several models which is difficult to understand in terms of source variability. During the Hanford PSHA (Hanford.gov, 2014) this was demonstrated to be an effect of sampling different ranges of site response from event to event. The site variability is, therefore, mapped into between-event terms leading to the peak at 0.1 s.

Arguments for observing lower variability at large magnitudes include the fact that meta-data for large events (e.g. magnitude, depth and mechanism) are more reliable. While this is, in general, true, there has been significant work in recent years to develop reliable earthquake catalogs for smaller events. Another argument is that, due to large earthquakes having large rupture sizes, the sensitivity of ground motion to, for example depth or magnitude, is less. For example,  $M < 5$  events can generally be assumed to be point sources, with amplitudes decaying in proportion to the reciprocal of hypocentral distance. On the other hand,  $M > 6$  events emit waves from a

667 range of sources along several kilometers of rupture. Increasing the depth or  
 668 size of this fault, whilst changing the distance over which some of the seismic  
 669 energy must propagate, will, therefore, have a reduced effect. This is evident  
 670 in the saturation of ground-motion amplitudes for increasing magnitude in  
 671 GMPEs. Having reliable meta-data for larger events is, therefore, arguably  
 672 less important than for small earthquakes for sites not close to major active  
 673 faults. For other locations, reliable information on fault geometry and other  
 674 properties (e.g. rupture mode) is vital when estimating near-source ground  
 675 motions.

676 The limited number of events at large magnitudes leaves  $\tau$  open to under-  
 677 sampling (with each event only contributing a single data-point to the esti-  
 678 mate of  $\tau$ ). Given that strong-motion databases often include only a handful  
 679 of well-recorded events with  $\mathbf{M} > 7$ , the reliability of heteroscedastic  $\tau$  can  
 680 be called into question. Comparing values from different GMPEs we can see  
 681 that the variability in  $\tau$  estimates is rather high (Figure 4). In reality,  $\tau$  is  
 682 likely to be heteroscedastic, but caution should clearly be applied in using  
 683 low values at  $\mathbf{M} > 7.5$  coming from extrapolation of trends from smaller  
 684 magnitudes (Musson, 2009). Models developed with constant  $\tau$  estimates  
 685 for  $\mathbf{M} < 5$  and  $\mathbf{M} > 7$  connected by a linear trend (e.g. Abrahamson et al.,  
 686 2014) are an appropriate compromise in this sense.

## 687 5.2. *Within-event variability*

688 Ground-motion variability with respect to a given GMPE for single event  
 689 is described by within-event variability ( $\phi$ ). It can be interpreted as describ-  
 690 ing the standard deviation of the misfit between GMPE and data after ac-  
 691 counting for the between-event terms. In terms of the random-effects frame-  
 692 work,  $\phi$  describes the standard deviation of within-event random-effects.

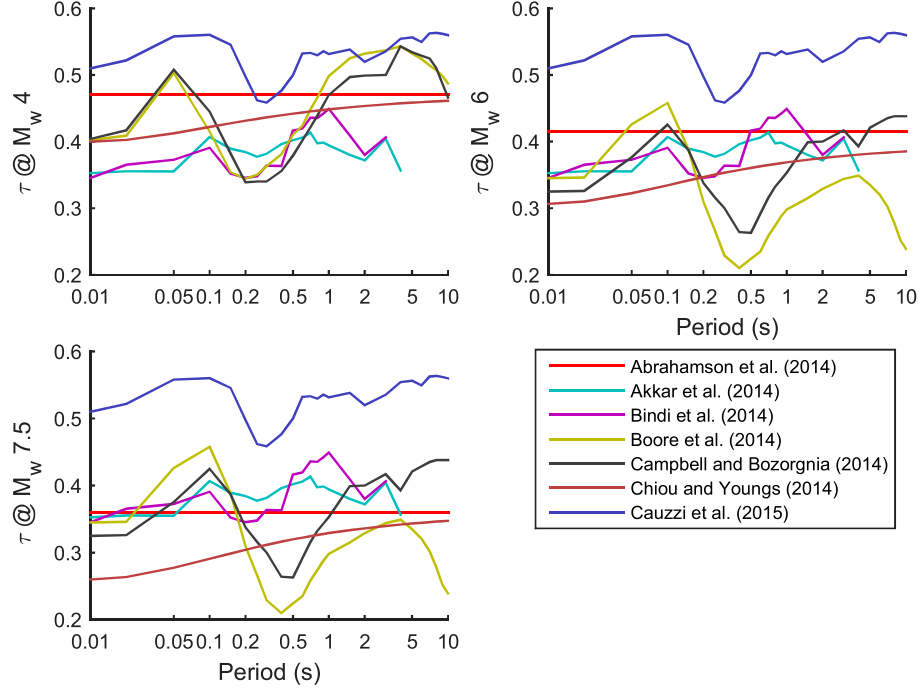


Figure 4: Comparison of the  $\tau$  models of six recent GMPs: Abrahamson et al. (2014), Boore et al. (2014), Campbell and Bozorgnia (2014) and Chiou and Youngs (2014) (predominantly Californian data); Bindi et al. (2014) and Akkar et al. (2014a) (European and the Middle Eastern data); and Cauzzi et al. (2015) (Japanese data), for  $M_w$  4, 6 and 7.5 with respect to response period.

693 The logarithm of ground-motion variability is assumed to be normally dis-  
 694 tributed. The total variability of a dataset with respect to a GMPE is then  
 695 given by (assuming independence between the two components):  $\sqrt{\tau^2 + \phi^2}$ .  
 696 Within-event variability is related to path and site phenomena in addition to  
 697 any spatially-dependent source characteristics, such as radiation pattern or  
 698 directivity effects. Because of the dominant effect of site amplification and  
 699 the significant variability of site effects these are considered to be a signif-  
 700 icant source of within-event variability (e.g. Rodriguez-Marek et al., 2011).  
 701 In the most recent studies,  $\phi$  is therefore split into components describing  
 702 site-to-site variability ( $\phi_{S2S}$ ) and within-site variability ( $\phi_0$ ). Drouet and  
 703 Cotton (2015) showed that the within-event variability is controlled by a  
 704 number of factors: the most significant being site amplification/attenuation  
 705 effects (including  $\kappa$ ) followed by path effects, such as geometrical and anelas-  
 706 tic attenuation. Bindi et al. (2014) observe that certain stations contribute  
 707 a large proportion of the soft soil (Eurocode 8 class D) sites for European  
 708 GMPEs. Some often-triggered stations, therefore, have strong influence on  
 709 the model and may reduce the apparent within-event variability.

710 While  $\phi$  is often considered a ‘site term’ it is also observed to be mag-  
 711 nitude, distance and  $V_{s,30}$  dependent (Figure 5). For instance, Boore et al.  
 712 (2014) and Campbell and Bozorgnia (2014) show that  $\phi$  decreases with mag-  
 713 nitude at short periods and increases with magnitude at long periods. Due  
 714 to the interaction of ergodic and non-ergodic components of variability it is  
 715 difficult to know if this is truly a site-specific effect or due to site-to-site vari-  
 716 ability (different sites having recorded different ranges of earthquake magni-  
 717 tudes and distances). An effective magnitude-distance dependence of  $\phi$  due  
 718 to nonlinearity of soil response has been incorporated into GMPE develop-  
 719 ment. For example, Abrahamson et al. (2014) account for soil non-linearity



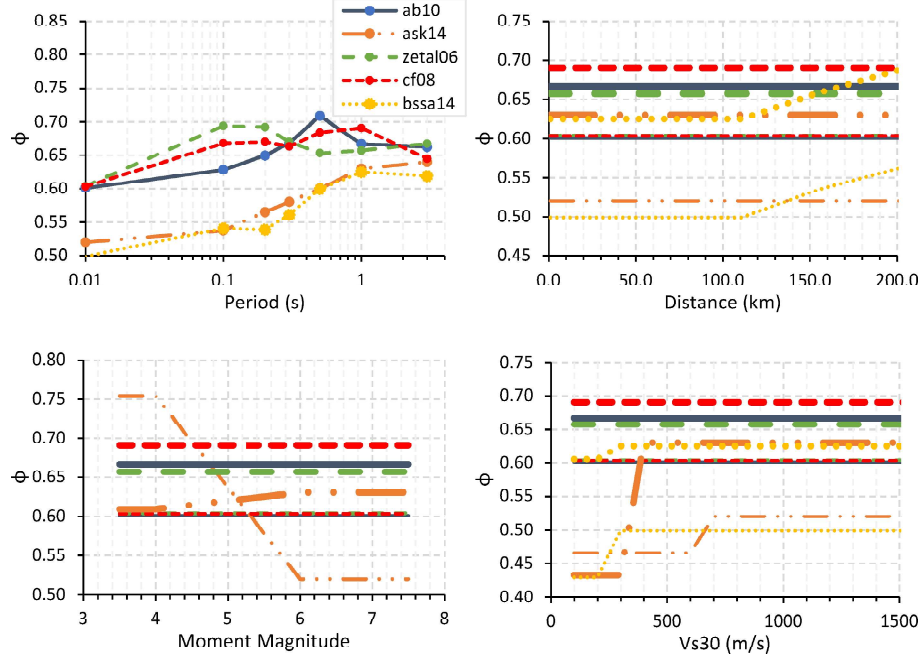


Figure 5: Comparison of estimates of the within-event variability  $\phi$  from some recent GMPEs, where ab10 corresponds to Akkar and Bommer (2010), ask14 corresponds to Abrahamson et al. (2014), zetal06 corresponds to Zhao et al. (2006), cf08 corresponds to Cauzzi and Faccioli (2008) and bssa14 corresponds to Boore et al. (2014).

reducing the variability of short-period motions. Focusing on non-ergodic sigma, Rodriguez-Marek et al. (2013) present models for single-station  $\phi$  using data from various tectonic regions. They show a decrease of single-station  $\phi$  over all periods, which differs from the observations of ergodic variability, where long-period motions show increased  $\phi$  for large earthquakes.

An explanation for the different observations of  $\phi$ 's dependency on distance and magnitude may be found in the dependence of response spectral amplification on the input motion (e.g. Bora et al., 2016). Given that resonance effects in site response depend greatly on the site type (e.g. long-period resonance for deep sedimentary basins and high-frequency resonance

730 for thin deposits of alluvium), whether or not input motions (broadly de-  
 731 fined by magnitude and distance) excite a particular resonant frequency will  
 732 make a difference to ground-motion variability. As a result, depending on  
 733 the characteristic site type(s) in a strong-motion database, the sensitivity  
 734 of  $\phi$  to magnitude and distance will vary. Rock, or hard-rock sites, will be  
 735 mostly independent of input motion, while soil and stiff-soil sites will be  
 736 strongly dependent on the input motions, with nearby smaller-magnitude  
 737 (higher-frequency) events strongly amplified by high-frequency resonance  
 738 peaks.

### 739 *5.3. Single-station variability*

740 The ergodic assumption has been used to derive most GMPEs to date  
 741 (Figure 6). This assumption is made to overcome the fact that limited data  
 742 are available at individual stations and to provide average (e.g. azimuth-  
 743 independent) predictions. The ergodic assumption assumes that spatial  
 744 variability can be mapped into variability in time (Anderson and Brune,  
 745 1999). Given that station-to-station variability is a significant component of  
 746 aleatory variability captured in GMPEs, this assumption cannot be valid for  
 747 a single site. To overcome this limitation, the concept of single-station vari-  
 748 ability was introduced by Anderson and Brune (1999) and first estimated  
 749 using a large set of data by Atkinson (2006).  $\sigma_{SS}$  describes the total vari-  
 750 ability (within- and between-event) in SA expected at a single site. Provided  
 751 ground-motion variability is separated into  $\phi_0$  and  $\phi_{S2S}$  then simply setting  
 752  $\phi_{S2S}$  to zero will result in  $\sigma_{SS}$ . Rodriguez-Marek et al. (2013) showed that  
 753  $\sigma_{SS}$  shows remarkably little variability between regions thereby suggesting  
 754 that it is the site-to-site variability that drives differences in ground-motion  
 755 variability between regions. Although recent work by Al Atik (2015) evi-

756 denced slightly higher values of  $\sigma_{SS}$  based on data from the stable continen-  
757 tal region of central and eastern North America.

758 While  $\sigma_{SS}$  reduces the variability to that consistent with what would  
759 be observed given sufficient recordings at a single site, we must be careful  
760 that the GMPE used for the single site is not biased. When GMPEs are  
761 derived using data from a variety of sites they invariably produce output  
762 that is consistent with the average site within a given site class or for a  
763 given  $V_{s,30}$  in the dataset.  $\phi_{S2S}$  then accounts for the variability between  
764 sites. However, if we are just looking at one site and using  $\sigma_{SS}$  we must  
765 ensure that the GMPE produces a median consistent with our study site.  
766 For this reason host-to-target adjustments (Section 4.5) may be used.

767 Building on current practice of using mixed-effects regression to deter-  
768 mine GMPE coefficients (Abrahamson and Youngs, 1992), Stafford (2014)  
769 presents the use of crossed and nested mixed effects to determine robust  
770 models that are not subject to the short comings of multi-stage approaches  
771 often adopted to separate model components. Using this approach he shows  
772 how site- and region-specific effects can be accounted for within a single  
773 inversion.

## 774 6. Epistemic uncertainty

775 Despite rapidly increasing strong-motion databases and the consider-  
776 able improvements in our understanding and modeling of strong ground  
777 motions (see above) each new GMPE published invariably predicts different  
778 levels of average shaking and its variability for every scenario than previ-  
779 ous models. These differences arise from epistemic uncertainty, although  
780 generally this uncertainty is larger than these differences imply. If we had

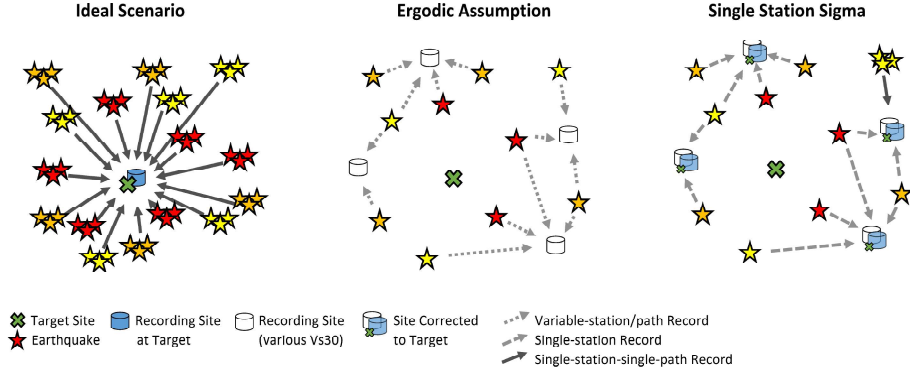


Figure 6: Sketch of transition from ergodic to (partial) non-ergodic assumption. Earthquakes of the same magnitude but with different characteristics (e.g. stress parameter) are indicated by different colored stars. Left: ideal scenario, with numerous events being recorded at a single station. Full separation of uncertainties related to event characteristics ( $\tau$ ), and path and site characteristics ( $\phi$ ) is possible down to single-event-single-path  $\sigma$ . Center: typical scenario, with events sparsely recorded on regional network with various site types (e.g.  $V_{s,30}$ ). An ergodic assumption is used: time equivalent to space to define  $\tau$  and  $\phi$ . Right: advanced approaches correct sites to account for differing response (single-site  $\sigma$ ), while multiple events on the same source (e.g. fault) allow single site-single-path  $\sigma$  to be defined.

781 an infinite amount of data available from every earthquake scenario, travel  
 782 path and site then the epistemic uncertainty would reduce to zero as there  
 783 would be no need for models, simply selection of the strong-motion records  
 784 from the database appropriate for the required scenario. There may still  
 785 be aleatory variability in this case because of intrinsic randomness in earth-  
 786 quake rupture, wave scattering and so forth but for a given scenario the  
 787 true average ground motions and its variability should be defined exactly.  
 788 Non-parametric methods (e.g. neural networks) are useful in investigating  
 789 ground-motion scaling for well-sampled scenarios (e.g. Derras et al., 2014;  
 790 Hermkes et al., 2014). Such data-mining approaches are likely to play an  
 791 increasing role as strong-motion databases grow.

792 The day of sufficient observations to no longer require models is many  
 793 decades, or even centuries, away for most scenarios of engineering interest.  
 794 As shown by Douglas (2010b, 2012) average predicted ground motions for  
 795 scenarios close to the barycenter of available data ( $M_w \sim 6$ ,  $R \sim 20$  km) have  
 796 remained roughly constant over the past few decades despite improvements  
 797 to GMPEs. For well-observed regions such as western North America there  
 798 has been some convergence in predictions (Douglas, 2010b). This is because  
 799 the same data are used to tune the models. Predictions for scenarios closer  
 800 to the edges of available observations (e.g.  $M_w > 7$  and  $R < 10$  km), how-  
 801 ever, display larger differences. One question that is rarely raised is: how  
 802 representative are the available data of ground motions in that region? For  
 803 example, are the few well-recorded  $M > 7$  crustal earthquakes in strong-  
 804 motion databases representative of all future large events? Re-sampling and  
 805 bootstrap techniques to assess the stability of the models to the removal of  
 806 data could be useful in this context (e.g. Berge-Thierry et al., 2003; Bindi  
 807 et al., 2014). These approaches, however, only provide guidance on the im-

808 pact of data that are already available and not on the stability of the models  
809 to *future* observations.

810 Another way of understanding epistemic uncertainties is to examine the  
811 statistical confidence limits (e.g. Draper and Smith, 1998) in the median  
812 predictions from a given GMPE (Campbell, 1985). This has been done  
813 by Douglas (2010a), who examined the width of the confidence limits from  
814 three generations of GMPEs for western North America (Joyner and Boore,  
815 1981; Boore et al., 1997; Boore and Atkinson, 2008) and Europe and the  
816 Middle East (Ambraseys and Bommer, 1991; Ambraseys et al., 1996, 2005).  
817 Douglas (2010a) finds that the confidence limits for the western North Amer-  
818 ican models are narrowing (and hence epistemic uncertainty is reducing) but  
819 that this is not seen for the models from Europe and the Middle East, which  
820 he relates to making the models too complex given the number of records  
821 available. Recently, Al Atik and Youngs (2014) compute confidence limits  
822 for the NGA West 2 GMPEs and propose a method to include this uncer-  
823 tainty within a seismic hazard assessment. A third way of examining simi-  
824 larities between models is to use high-dimensional information-visualization  
825 techniques, such as Sammon’s maps (Scherbaum et al., 2010), that display  
826 models on a 2D graph thereby allowing identification of models that predict  
827 similar motions.

828 As strong-motion networks become denser the average number of sta-  
829 tions that record a given earthquake increases, which means that model  
830 source terms (e.g. style-of-faulting factors) and the between-event variabil-  
831 ity ( $\tau$ ) are better constrained in recent GMPEs. Similarly a modern station  
832 generally records more earthquakes leading to better estimates of site terms  
833 and single-station  $\sigma$ . Site terms are now less biased since fewer stations con-  
834 tribute a large proportion of records to strong-motion databases, although

835 the number of records per station remains highly variable.

836 The reduction of epistemic uncertainty (differences in predictions among  
837 models) remains a considerable challenge. It is vital that this uncertainty  
838 is not artificially reduced but that seismic hazard assessments correctly ac-  
839 count for the true uncertainty in ground-motion prediction. There is a  
840 trade-off to be made between including more and more independent vari-  
841 ables to seek to reduce  $\sigma$  but thereby increasing epistemic uncertainty in  
842 the model because these variables are difficult to predict before an earth-  
843 quake and because more variables require more data to constrain the free  
844 coefficients in the GMPE.

845 Only a few GMPE developers (e.g. Douglas et al., 2013) estimate the  
846 epistemic uncertainty in their models. Estimates of the lower bound of the  
847 epistemic uncertainty can be made by comparing multiple models by the  
848 same developer team or by various teams using the same master database  
849 (Douglas et al., 2014a; Abrahamson et al., 2008; Gregor et al., 2014). These  
850 comparisons do not capture the part of uncertainty related to the question:  
851 for which parts of the models are changes likely in the future because of lack  
852 of understanding or knowledge? The motto of US General Colin Powell:  
853 ‘Tell me what you know. Tell me what you don’t know. Then tell me  
854 what you think. Always distinguish which is which’ may be useful in this  
855 context. The first and third parts of this saying are remembered by all  
856 GMPE developers but the second and last parts are often forgotten in the  
857 development of ground-motion models.

858 Logic trees (Kulkarni et al., 1984) are used within seismic hazard assess-  
859 ment to model epistemic uncertainty by assigning weights to each ground-  
860 motion model, for example, depending on the degree of belief that the haz-  
861 ard analyst has in that model being the appropriate one for the study (e.g.

862 Bommer et al., 2005). Consequently there should be a correlation between  
863 the level of understanding about earthquake shaking at the study site (or  
864 regions) and the spread in predicted median ground motions from the logic  
865 tree: wider spread in predictions where knowledge is limited and reinforcing  
866 predictions where knowledge is greater. There is, however, evidence  
867 for ‘group think’ in models. For example, many of the predictions from  
868 the NGA models changed in the same way from 2008 (NGA West 1) to  
869 2014 (NGA West 2), e.g. the predictions for earthquakes with  $M < 5.5$   
870 change considerably [and in agreement with what would be expected (Bom-  
871 mer et al., 2007)] but those for  $M7.5$  change very little (Figure 7). Will such  
872 large changes to predictions also occur when more large earthquakes have  
873 been well recorded? When there are few observations it is uncomfortable  
874 to be out on a limb and for your model to predict greatly different motions  
875 than the majority of models. Consequently, things have changed where new  
876 data (e.g. small magnitudes) are added to strong-motion databases but not  
877 where uncertainty remains high, e.g. close to large events. This leads to  
878 the apparently inconsistent observation made by Douglas (2010b) that the  
879 divergence in predictions of median ground motions from GMPEs for stable  
880 continental regions is lower for large magnitudes (for which there are very  
881 few observations) than for small magnitudes (where data exist).

882 Since about 2010 there has been increasing use of the backbone approach  
883 (Atkinson et al., 2014) to model epistemic uncertainty in ground-motion  
884 prediction. In this approach, rather than use a suite of GMPEs to model  
885 epistemic uncertainty within a logic tree, a single GMPE (or sometimes two  
886 or three GMPEs) is scaled up and down by factors to generate a set of  
887 mutually-exclusive and collectively-exhaustive models. The backbone ap-  
888 proach has the advantage of always having an overall ground-motion model



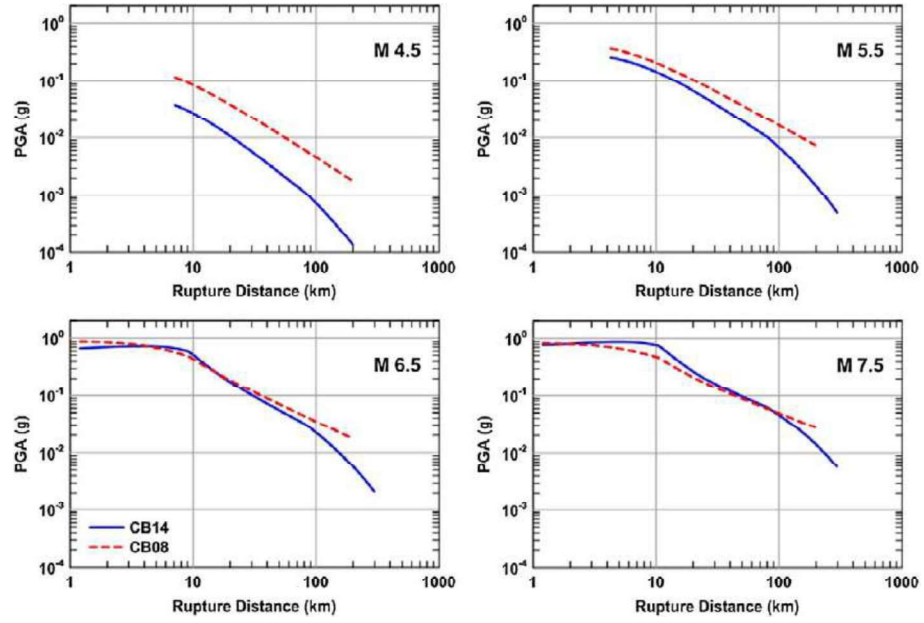


Figure 7: Comparison of predicted median PGA from Campbell and Bozorgnia (2008) (CB08) and Campbell and Bozorgnia (2014) (CB14) on a site with  $V_{s,30} = 760$  m/s for M4.5 to 7.5 from 45°-dipping reverse fault. Figure taken from Campbell and Bozorgnia (2014).

that allows the epistemic uncertainty to be defined directly by expert judgment, and which is explicitly definable. The multiple GMPEs approach, however, leads to varying modeled uncertainties, which may lead to pinch points for certain scenarios that may not be logical (e.g. where there are few data but the GMPEs coincide). The backbone approach, however, may lead to overestimation of epistemic uncertainties when data are abundant and it can be tricky to calibrate. On the other hand, the availability of abundant data is unfortunately not presently the case for all relevant scenarios (e.g. large magnitude near-source) and using only published GMPEs without any scaling factors will likely lead to underestimation of the true epistemic uncertainty.

## 7. Extensions to ground-motion models

As noted above, the vast majority of GMPEs have been derived for PGA and linear elastic response spectral ordinates (particularly for 5% of critical damping). Because of its proposed use in liquefaction analysis, its better correlation with felt and damage reports and its use in some regulations (e.g. Bommer and Alarcón, 2006) PGV has also become a popular IM for ground-motion models. In the past decade or so, there has been a growing interest in deriving models for other IMs (Douglas, 2012), in particular Arias intensity (Arias, 1970) [commonly used in the analysis of earthquake-triggered landslides (e.g. Harp and Wilson, 1995)], relative significant duration (Trifunac and Brady, 1975) and peak ground displacement. A handful of models for other IMs (e.g. Fourier spectral amplitudes, Japanese Meteorological Agency seismic intensity, cumulative absolute velocity, mean spectral period and inelastic response spectral ordinates) have also been published (Douglas,

914 2016). Finally, there is a growing set of macroseismic intensity prediction  
915 equations (Cua et al., 2010). These allow PSHA to be conducted directly  
916 for IMs that have various engineering uses rather than having to conduct  
917 a seismic hazard assessment for PGA, for example, and then convert this  
918 to the required IM. This should lead to smaller overall uncertainties within  
919 risk assessments.

920 Standard GMPEs predict independent scalar IMs. This is what is re-  
921 quired by PSHA to compute uniform hazard spectra, for example. Re-  
922 cent developments in earthquake engineering, e.g. conditional mean spectra  
923 (Baker, 2011), mean that it is important to know the correlation between  
924 spectral ordinates at different structural periods (e.g. Baker and Jayaram,  
925 2008) and between various IMs (e.g. Bradley, 2011). Consequently models  
926 for the estimation of these correlations have been derived. These provide a  
927 more complete assessment of earthquake ground motions.

928 Another way in which the picture of earthquake shaking is becoming  
929 richer is through the derivation of models to estimate the spatial correlation  
930 of motions between neighboring geographical locations (e.g. Goda and Hong,  
931 2008). Such models improve the accuracy of earthquake loss predictions of  
932 spatially-distributed portfolios (e.g. Weatherill et al., 2015).

## 933 **8. Conclusions and ways forward for ground-motion prediction**

934 A number of multinational projects have, over the last decade, brought  
935 significant advances in ground motion characterization for seismic hazard  
936 analyses. These include the NGA West 1 and 2 (Power et al., 2008; Bo-  
937 zorgnia et al., 2014), NGA East (Pacific Earthquake Engineering Research  
938 Center, 2015) and RESORCE (Akkar et al., 2014b) projects. In addition to

939 these initiatives, numerous peer-reviewed articles have improved our knowl-  
 940 edge and understanding of ground-motion prediction in a variety of regions,  
 941 from active regions with high seismicity (mainly empirical GMPEs) to sta-  
 942 ble continental regions with low seismicity (with focus on robust simula-  
 943 tion approaches, such as stochastic methods). Despite the significant in-  
 944 vestment over the last decades, the aleatory variability in ground-motion  
 945 prediction for scenario events appears not to have decreased (e.g. Strasser  
 946 et al., 2009). Nevertheless, our understanding of the source and behavior  
 947 of ground-motion variability has improved dramatically, with articles barely  
 948 mentioning it 20 years ago, to the current state where sometimes roughly  
 949 half of a manuscript presenting a new GMPE is dedicated to its charac-  
 950 terization. While the total variability is therefore not reduced, the way in  
 951 which it is implemented in hazard models is now more realistic. The biggest  
 952 improvement is arguably the shift from ergodic towards non-ergodic variabil-  
 953 ity. This has reduced the  $\sigma$  used within site-specific (or reference-specific)  
 954 hazard analyses by as much as 30%.

955 Despite the great advances of recent years, ground-motion characteriza-  
 956 tion is still very much a topic in development. Some authors (e.g. Atkinson,  
 957 2012) have predicted that the goal is for numerical simulations to be per-  
 958 formed to estimate ground motion and its variability. Despite the increase  
 959 in computing power allowing the calculation of shorter-period ground mo-  
 960 tions (with current limits around 0.3 to 1 s), the limitation of simulations is  
 961 twofold. Firstly, they rely on geophysical characterization of the crust and  
 962 shallow subsurface, but at short-periods ( $< 1$  s) the resolution scale of most  
 963 available geophysical models is simply insufficient. To overcome this lim-  
 964 itation, so-called hybrid approaches are used, where stochastic simulation  
 965 models are implemented to some cross-over period (e.g. Graves and Pitarka,

2010). Such methods clearly have the same limitations of existing empirical and stochastic models at short periods. Purely deterministic numerical simulations are still, therefore, at least several years away. The second limitation of numerical simulations is the understanding of constituent parameters and their covariances. Engineering practice requires stable and repeatable models, which GMPEs provide. While numerical simulations can be calibrated to provide predictions consistent with observed earthquake shaking, in practice the input parameters are poorly understood meaning that naive simulations may be incorrect.

Before purely deterministic numerical scenario-simulations become possible the most promising developments in PSHA lie with the understanding of ground-motion variability, which drives hazard at long return-periods. The conceptual approach of single-station (non-ergodic) sigma provides the framework for this. However, most datasets are still significantly lacking in data where they are of most relevance for long return-period hazard (records in the upper tails of the ground-motion distribution from moderate earthquakes and large events recorded at near distances). The robustness of models describing this variability is, therefore, called into question. Improved approaches for modeling data with mixed sampling in the model space, obtaining additional empirical data, and the reliable simulation of such data is, therefore, of great importance.

In some senses, seismology is analogous to economics in that we cannot do full-scale controlled experiments, e.g. we cannot replay an earthquake (seismology) or a recession (economics) with slightly altered input parameters. Unlike economics, however, in seismology we generally do not have masses of data. Perhaps there are some statistical tools and approaches that are used in economics that could be applied to seismological data or models,

993 e.g. in the assessment of epistemic uncertainty. Although as noted by, for  
994 example, Kahneman (2012) experts in economics and in other fields find  
995 it challenging to correctly assess what they know and, equally important,  
996 what they do not know. There is clearly a need in ground-motion prediction  
997 to improve the calibration of the level of epistemic uncertainty modeled by  
998 GMPEs within seismic hazard assessments.

999 Douglas et al. (2014b) find that often the more expensive, carefully-  
1000 undertaken assessments for single sites model *higher* uncertainty than cheaper  
1001 regional assessments, which is a demonstration of an inconsistency in cap-  
1002 turing epistemic uncertainty. However, it should be noted that the primary  
1003 objective of more elaborate assessments, such as those following the SSHAC  
1004 guidelines (Budnitz et al., 1997), is to ensure the capture of epistemic uncer-  
1005 tainty. The higher study levels in SSHAC increase the likelihood of this ob-  
1006 jective being met. Therefore, it should not surprise us that the uncertainty  
1007 ranges from SSHAC Level 3 or 4 studies are greater than those in small  
1008 studies performed more informally by an individual or a small team. On  
1009 the other hand, epistemic uncertainty is reduced by data collection. In the  
1010 Thyspunt PSHA (Bommer et al., 2015), for example, without the historical  
1011 seismicity studies, geological investigations and extensive velocity measure-  
1012 ments at the site, the total uncertainty in the final hazard assessments would  
1013 have been considerably larger. More expensive studies are, therefore, forced  
1014 to undertake more analyses to assure that epistemic uncertainty is reduced,  
1015 as opposed to smaller studies that may simply make an assumption that the  
1016 overall epistemic uncertainty is at a given level.

1017 The growth of unconventional gas and oil extraction and associated fluid  
1018 injection and, to a lesser extent, geothermal energy has led to a significant  
1019 increase in induced seismicity (Rubinstein and Mahani, 2015). This fo-

cus has seen several GMPEs being published for the purpose of predicting ground motion from small earthquakes at very short distances. While common wisdom would suggest that damage due to induced seismicity, which is generally limited to events with  $M < 5$ , is negligible, there have been cases of significant insured losses (Giardini, 2009), although what proportion of damage is earthquake-related is debatable.

As noted above, some authors (Field et al., 2003; Atkinson, 2012) have argued that GMPEs will soon be replaced by numerical simulations of earthquake shaking. Such simulations do provide a much richer representation of the earthquake hazard to engineers (full time-histories rather than simply intensity measures) and they allow source- and site-specific calculations, although for a limited structural period range. For poorly-sampled magnitude-distance ranges and unusual source (e.g. deep crustal sources), path (e.g. strong velocity contrasts) and site conditions (e.g. nonlinear soils) simulations are invaluable in guiding the development of GMPEs. The general consensus is that full-waveform simulation approaches are currently not sufficiently constrained, however, to form the basis of hazard analyses due to their reliance on a full understanding of the physical system (including effects such as plastic deformation, fault shape and roughness). They are at a stage, however, where simulations provide valuable insight into the expected behavior of source effects and wave propagation in heterogeneous media, which can be combined with empirical data and analyses. Although ground-motion simulations show significant advances with the advent of high-performance computing and the development of better procedures, GMPEs are likely to remain a key component of hazard assessments for the foreseeable future.

One attractive approach to ground-motion simulation is ‘virtual earthquakes’ (Denolle et al., 2014), in which the Green’s functions measuring the

1047 Earth’s response to point impulses are derived from the ambient seismic field  
1048 (i.e. microtremors) and then these are used to predict ground motion from  
1049 a series of point sources to model fault rupture. This approach captures the  
1050 effect of travel path in the region, e.g. sedimentary basin effects, but it is  
1051 currently restricted to structural periods longer than 3 s. For long periods  
1052 it may be possible to simulate ground motions using this technique for the  
1053 derivation of ground-motion models but an outstanding issue is assessing  
1054 the variability and uncertainty associated with these simulations.

1055 Treverton (2007) discusses the difference between a puzzle and a mys-  
1056 tery. To solve a puzzle you need more information while to solve a mystery  
1057 requires clever analysis of the information that is already available. Ground-  
1058 motion prediction currently is more of a puzzle, because data are limited,  
1059 whilst it is often seen as a mystery, where complex analysis is applied to  
1060 very little data. As noted by Atkinson (2004) for ‘every dollar that is spent  
1061 trying to quantify uncertainty, we should spend 10 dollars collecting and an-  
1062 alyzing data that would reduce uncertainty’. While we have seen significant  
1063 changes in many, if not most, recent PSHAs compared to earlier studies,  
1064 due to the advancement of state-of-practice, a significant contribution to  
1065 this can be put down to the availability of new data and better treatment of  
1066 it in PSHA. Collection of more strong-motion data and, equally important,  
1067 the associated metadata (e.g. local site conditions) is the only reliable way  
1068 of reducing uncertainty in ground-motion prediction and hence it should be  
1069 prioritized. With the rapid decrease in the cost of strong-motion instrumen-  
1070 tation and the ease-of-use of new sensors, there is hope that the era of only  
1071 recording a single near-source accelerogram from a **M**7.8 earthquake [as was  
1072 the case for the Gorkha (Nepal) earthquake of 25th April 2015] is coming  
1073 to an end. Strong-motion monitoring in seismic areas could be encouraged



1074 by, for example: providing instruments to schools for use as an educational  
1075 tool, installing sensors in public buildings, and requiring instrumentation  
1076 as part of the building code for infrastructure (e.g. power plants). Large  
1077 earthquakes occur infrequently and they present an opportunity to signifi-  
1078 cantly improve our knowledge of earthquake shaking, which is vital in the  
1079 reduction of seismic risk.

1080 Our understanding of earthquake hazard has improved dramatically in  
1081 the past decades. Therefore, is it necessary to continue refining seismic  
1082 hazard assessments when the results are unlikely to change dramatically?  
1083 We argue that such refinement is required if not from a purely scientific  
1084 point of view but because it is important from the regulator’s viewpoint  
1085 that all avenues are explored and the best analysis is performed. Many drug  
1086 trials are conducted that demonstrate that a drug is not useful but it is  
1087 not then argued that the trial was a waste of money – why should seismic  
1088 hazard assessment be any different? The seismological community cannot  
1089 be seen to be resting on our laurels and not striving for improved knowledge  
1090 and understanding. In addition, while significant recent advances have been  
1091 made in education, it is necessary to continue to train the next generation  
1092 of engineering seismologists so that they can produce high-quality hazard  
1093 assessments and, equally important, to understand what such assessments  
1094 mean. Examples of this should focus on two important elements: a) hands-  
1095 on experience in real projects (most training is typically theoretical and in  
1096 the authors’ experience is not completely aligned with real projects), and b)  
1097 funding science and data collection underlying earthquake engineering and  
1098 engineering seismology.

1099 Finally, while significant advances have been made in ground-motion  
1100 prediction over the past decade, we are continually surprised by unexpected

1101 events. Recent examples include the high PGAs recorded during the **M9**  
 1102 Tohoku earthquake (2.7 g); the long-period (3-5 s) motions (over 4 m/s)  
 1103 recorded during the **M7.8** Gorkha, Nepal event with recorded peak displace-  
 1104 ments of up to 1.87 m; and in lower seismicity areas the Market Rasen (**M4.5**,  
 1105 UK) and St Die (**M4.8**, France) earthquakes (Ottemöller and Sargeant, 2010;  
 1106 Scherbaum et al., 2004), which exhibited much higher than expected motions  
 1107 than expected using local ground-motion models. It is clear, therefore, that  
 1108 while advances are welcome in aspects such as median predictions and the  
 1109 capture of uncertainty, we still lack full understanding of the fundamentals  
 1110 of source-, path- and site-specific earthquake ground motion.

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 1114 their careful reviews of an early version of this article, which led to significant  
 1115 improvements to the article.

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